

#39 $Y \sim \text{Gamma}(2, \beta)$

Note that, $U = \frac{2Y}{\beta} \sim \chi^2(4)$ is a pivot

$$P(0.71072 \leq \frac{2Y}{\beta} \leq 9.48773) = 0.90$$

Therefore $(\frac{2Y}{9.48773}, \frac{2Y}{0.71072})$ forms a 90% C.I. for β .

#41b $Y \sim N(0, \sigma^2)$,

Note that $\frac{Y^2}{\sigma^2} \sim \chi^2(1)$ is a pivot.

$$P(0.0039321 \leq \frac{Y^2}{\sigma^2}) = 0.95$$

$$\Rightarrow P(\sigma^2 \leq \frac{Y^2}{0.0039321}) = 0.95$$

$\Rightarrow \frac{Y^2}{0.0039321}$ is the 95% upper C.B. for σ^2 .

#42b $P(\frac{Y^2}{\sigma^2} \geq 0.0039321) = 0.95$

$$\Rightarrow P(\sigma \leq \sqrt{\frac{Y^2}{0.0039321}}) = 0.95$$

$\Rightarrow \frac{Y}{0.0627}$ is the 95% upper C.B. for σ

#44 $f_Y = \frac{2(\theta - Y)}{\theta^2}$ for $0 < Y < \theta$

(b) show that $\frac{Y}{\theta}$ is a pivotal quantity

Pf. transformation method:

let $U = \frac{Y}{\theta}$

step 1: $Y = \theta U$

step 2: θ

step 3: $f_U = \frac{2(\theta - \theta U)}{\theta^2} \cdot \theta = \frac{2\theta(1-U)\theta}{\theta^2} = 2(1-U)$

$U \sim \text{Beta}(\alpha=1, \beta=2)$ for $0 \leq U \leq 1$

Therefore, U is a pivotal quantity as it doesn't depend on any unknown parameters

(c) $P(U \leq a) = 0.9$

that is $\int_0^a 2(1-u) du = 0.9 \Rightarrow a = 1 - \sqrt{0.1} = 0.6838$

$\Rightarrow P(U \leq 0.6838) = P\left(\frac{Y}{\theta} \leq 0.6838\right) = P\left(\theta \geq \frac{Y}{0.6838}\right) = 0.9$

Therefore, the 90% Lower Bound for θ is $\left(\frac{Y}{0.6838}\right)$

#48

 $y_1, y_2, \dots, y_n \stackrel{iid}{\sim} \text{Gamma}(\alpha=2, \beta)$ where β is unknown.

(a) Let $U = \frac{2 \sum_{i=1}^n y_i}{\beta}$

Recall: if $U = a_1 X_1 + a_2 X_2 + \dots + a_n X_n$

Then $M_U(t) = M_{X_1}(a_1 t) \cdot M_{X_2}(a_2 t) \cdots M_{X_n}(a_n t)$

Here, $a_1 = a_2 = \dots = a_n = \frac{2}{\beta}$

Therefore, $M_U(t) = M_{Y_1}(\frac{2}{\beta}t) \cdot M_{Y_2}(\frac{2}{\beta}t) \cdots M_{Y_n}(\frac{2}{\beta}t)$

where $M_{Y_i}(t) = \left(\frac{1}{1-\beta t} \right)^2$

So, $M_U(t) = \left(\frac{1}{1-\frac{2}{\beta}\beta t} \right)^{2n} = \left(\frac{1}{1-2t} \right)^{2n}$

This is the MGF for the $\chi^2(4n)$

Clearly, $U = \frac{2 \sum_{i=1}^n y_i}{\beta}$ is a pivot.

(b) $P(a \leq U \leq b) = 0.95$ where $U \sim \chi^2(4n)$

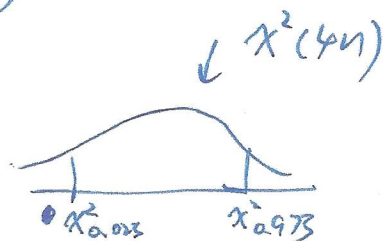
$\begin{cases} P(U \leq a) = 0.025 \\ P(U \leq b) = 0.975 \end{cases}$

$\Rightarrow P(\chi^2_{0.975}(4n) \leq U \leq \chi^2_{0.025}(4n)) = 0.95$

$\Rightarrow P(\chi^2_{0.975}(4n) \leq \frac{2 \sum_{i=1}^n y_i}{\beta} \leq \chi^2_{0.025}(4n)) = 0.95$

Therefore, the 95% C.I. for β is $\left(\frac{2 \sum_{i=1}^n y_i}{\chi^2_{0.975}}, \frac{2 \sum_{i=1}^n y_i}{\chi^2_{0.025}} \right)$

(c) C.I. is $\left(\frac{(2)(5)(1.39)}{34.1696}, \frac{(2)(5)(5.39)}{9.59083} \right) \text{ or } (1.577, 5.620)$



Supplementary Exercise

$$(1) f(x|\theta) = 1 \text{ for } \theta - \frac{1}{2} < x < \theta + \frac{1}{2}$$

step 1:
let $u = x - \theta$

$$(1) x = u + \theta$$

$$(2) 1$$

$$(3) f(u) = 1 \text{ for } -\frac{1}{2} \leq u < \frac{1}{2}$$

thus, $u = x - \theta$ is a pivot

step 2:

$$p(a < u < b) = 1 - \alpha$$

that's
$$\begin{cases} \int_a^b 1 du = 1 - \alpha \\ \int_{-\frac{1}{2}}^a 1 du = \frac{\alpha}{2} \end{cases} \Rightarrow \begin{cases} b - a = 1 - \alpha \\ a + \frac{1}{2} = \frac{\alpha}{2} \end{cases} \Rightarrow \begin{cases} a = \frac{\alpha}{2} - \frac{1}{2} \\ b = \frac{1}{2} - \frac{\alpha}{2} \end{cases}$$

$$\text{Thus, } p\left(\frac{\alpha}{2} - \frac{1}{2} < u < \frac{1}{2} - \frac{\alpha}{2}\right) = 1 - \alpha$$

step 3: $p\left(\frac{\alpha}{2} - \frac{1}{2} < x - \theta < \frac{1}{2} - \frac{\alpha}{2}\right) = 1 - \alpha$

$$p\left(x - \frac{1}{2} + \frac{\alpha}{2} < \theta < x - \frac{\alpha}{2} + \frac{1}{2}\right) = 1 - \alpha$$

The $100(1-\alpha)\%$ C.I for θ is $\left(x - \frac{1}{2} + \frac{\alpha}{2}, x - \frac{\alpha}{2} + \frac{1}{2}\right)$

$$(2) f(x|\theta) = \frac{2x}{\theta^2} \quad \text{for } 0 < x < \theta \text{ and } \theta > 0$$

Step 1: let $u = \frac{x}{\theta}$.

① $x = u\theta$

② θ

③ $f_u(u) = \frac{2 \cdot u\theta}{\theta^2} \cdot \theta = 2u \quad \text{for } 0 < u < 1$

Thus, $u = \frac{x}{\theta}$ is a pivot

Step 2: $P(a < u < b) = 1 - \alpha$

That is
$$\begin{cases} \int_a^b 2u \, du = 1 - \alpha \\ \int_0^a 2u \, du = \frac{\alpha}{2} \end{cases} \Rightarrow \begin{cases} b^2 - a^2 = 1 - \alpha \\ a^2 = \frac{\alpha}{2} \end{cases} \Rightarrow \begin{cases} b = \sqrt{1 - \frac{\alpha}{2}} \\ a = \sqrt{\frac{\alpha}{2}} \end{cases}$$

Thus, $P\left(\sqrt{\frac{\alpha}{2}} < \frac{x}{\theta} < \sqrt{1 - \frac{\alpha}{2}}\right) = 1 - \alpha$

Step 3: $P\left(\frac{x}{\sqrt{1 - \frac{\alpha}{2}}} < \theta < \frac{x}{\sqrt{\frac{\alpha}{2}}}\right) = 1 - \alpha$.

The $100(1 - \alpha)\%$ C.I. for θ is $\left(\frac{x}{\sqrt{1 - \frac{\alpha}{2}}}, \frac{x}{\sqrt{\frac{\alpha}{2}}}\right)$.

M440 - HW #4 - Solution

Section 8.6 problems:

#56

(a) with $Z_{0.01} = 2.326$, the 98% C.I. is $0.45 \pm 2.326 \sqrt{\frac{(0.45)(0.55)}{800}}$

(or) 0.45 ± 0.041 (or) $(0.409, 0.491)$

(b) since the value 0.50 is not contained in the interval, there is not compelling evidence that a majority of adults in Nov, 2003 are baseball fans.

#58 The parameter of interest is μ = mean number of days required for treatment. The 95% C.I. is approximately $\bar{y} \pm Z_{0.025} (\frac{s}{\sqrt{n}})$

(or) $5.4 \pm 1.96 (\frac{3.1}{\sqrt{500}})$ (or) $(5.13, 5.67)$

#61 with $Z_{0.025} = 1.96$, the 95% C.I. is $167.1 - 140.9 \pm 1.96 \sqrt{\frac{(24.3)^2 + (17.6)^2}{30}}$

(or) $(15.46, 36.94)$

#65 (a) The 98% C.I. is, with $Z_{0.01} = 2.326$ is

$$0.18 - 0.12 \pm 2.326 \sqrt{\frac{(0.18)(0.82) + (0.12)(0.88)}{100}}$$

(or) 0.06 ± 0.117 (or) $(-0.057, 0.177)$

(supplementary Ex 1):

$$(a) \quad E(\bar{X}) = E\left(\frac{1}{n} \sum X_i\right) = \frac{1}{n} (n\lambda) = \lambda$$

(b) Recall: For large n , we have $\hat{\theta} \pm Z_{\alpha/2} \cdot \sigma_{\hat{\theta}}$

$$\hat{\theta} = \bar{X} = \hat{\lambda}$$

$$\sigma_{\hat{\theta}} = \sqrt{V(\bar{X})} = \sqrt{\frac{\lambda}{n}} \approx \frac{\hat{X}}{n} = \frac{\hat{\lambda}}{n}$$

$$\Rightarrow \hat{\lambda} \pm Z_{0.025} \cdot \sqrt{\frac{\hat{X}}{n}} \Rightarrow \bar{X} \pm 1.96 \sqrt{\frac{\bar{X}}{n}}$$

$$(c) \quad n=30, \quad \bar{X}=10$$

$$\Rightarrow 95\% \text{ C.I. for } \lambda: \left(10 - 1.96 \sqrt{\frac{10}{36}}, 10 + 1.96 \sqrt{\frac{10}{36}} \right) \\ = (8.967, 11.033)$$