

Math 4200 HW #1 Solution

Section 6.3 problems:

#2 $f(y) = \begin{cases} (3/2)y^2 & -1 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases}$

The distribution function of Y is: $F_Y(y) = \int_{-1}^y (\frac{3}{2})t^2 dt = (\frac{1}{2})(y^3 - (-1)^3)$
for $-1 \leq y \leq 1$.

(a) $U = 3Y$,

$$F_U(u) = P(U \leq u) = P(3Y \leq u) = P(Y \leq \frac{u}{3}) = F_Y(\frac{u}{3}) = (\frac{1}{2})\left(\frac{u^3}{27} - (-1)^3\right)$$

for $-3 \leq u \leq 3$,

$$\Rightarrow f_U(u) = (F_U(u))' = \frac{u^2}{18}, \text{ for } -3 \leq u \leq 3$$

#33 $f(y) = \begin{cases} (\frac{3}{2})y^2 + y & 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases}$

$$U = 5 - (Y/2)$$

step 1: $g^{-1}(u) = Y = 2(5 - u) = 10 - 2u$

step 2: $\frac{dg^{-1}(u)}{du} = -2$

step 3: $f_U(u) = \left[\left(\frac{3}{2}\right)(10 - 2u)^2 + (10 - 2u) \right] |-2|$

$$= 4(80 - 31u + 3u^2) \text{ for } 4.5 \leq u \leq 5$$

#46

$$Y \sim \text{Gamma}\left(\frac{n}{2}, \beta\right), \quad W = \frac{2Y}{\beta}$$

show $W \sim \chi^2(n)$

$$= E[e^{\left(\frac{2t}{\beta}\right)Y}]$$

Pf. $M_W(t) = E[e^{Wt}] = E[e^{2Yt/\beta}] = E[e^{\left(\frac{2t}{\beta}\right)Y}] = M_Y\left(\frac{2t}{\beta}\right)$

since $Y \sim \text{Gamma}\left(\frac{n}{2}, \beta\right)$,

then $E_Y(t) = \left(\frac{1}{1-\beta t}\right)^{\frac{n}{2}}$

$$\Rightarrow M_W(t) = M_Y\left(\frac{2t}{\beta}\right) = \left(\frac{1}{1-\beta \cdot \frac{2t}{\beta}}\right)^{\frac{n}{2}} = \left(\frac{1}{1-2t}\right)^{\frac{n}{2}}$$

$$\Rightarrow W \sim \chi^2(n)$$

#6.74 Since $y_i \in \text{unif}(0, \theta)$, then $f_{Y_i}(y_i) = \frac{1}{\theta}$ for $0 \leq y_i \leq \theta$.

and $F_{Y_i}(y_i) = \frac{y}{\theta}$ for $0 \leq y \leq \theta$.

(a). $F_{Y_{(n)}}(y) = \left(\frac{y}{\theta}\right)^n$, for $0 \leq y \leq \theta$

(b) $f_{Y_{(n)}}(y) = ny^{n-1}/\theta^n$, for $0 \leq y \leq \theta$

(c). $E(Y_{(n)}) = \int_0^\theta (y) (ny^{n-1}/\theta^n) dy = \left(\frac{n}{n+1}\right)\theta$

$$V(Y_{(n)}) = \frac{n\theta^2}{(n+1)^2(n+2)}$$

$\uparrow$ $n \text{th } E(Y_{(n)}^2) - (E(Y_{(n)}))^2$

(#1) $Z \sim N(0,1)$, let $U = Z^2$, show $U \sim \chi^2(1)$

Pf:

Use MGF method:

$$\begin{aligned} M_U(t) &= E[e^{ut}] = E[e^{Z^2 t}] = \int_{-\infty}^{\infty} e^{z^2 t} \cdot f_Z(z) dz \\ &= \int_{-\infty}^{\infty} e^{z^2 t} \cdot \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz \\ &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{(t - \frac{1}{2})z^2} dz \quad \dots (*) \end{aligned}$$

Recall: If $X \sim N(0, \sigma^2)$, then $f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2}$

$$\text{So, } \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2} dx = 1$$

$$\Leftrightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx = \sigma \quad (**)$$

Rewrite (*), we have:

$$(*) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{-z^2(1-2t)}{2}} dz = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{-z^2}{2(1-2t)}} dz$$

$$\stackrel{\text{According to } (**), \text{ Here } \sigma^2 = \frac{1}{1-2t}}{\uparrow} \left(\frac{1}{1-2t} \right)^{\frac{1}{2}} \text{ is MGF for } \chi^2(1)$$

Therefore, $U = Z^2 \sim \chi^2(1)$.

(#2) $M_Y(t) = E(e^{Yt}) = E(e^{\frac{2\alpha t}{\beta}}) = M_X\left(\frac{2t}{\beta}\right) = \left[\frac{1}{1 - \beta \cdot (\frac{2t}{\beta})} \right]^\alpha = \left(\frac{1}{1-2t} \right)^\alpha$
 this is the MGF for $\chi^2(2\alpha)$.

#3 Let $u \sim \text{unif}(0,1)$, Define $X = -\log u$ and $Y = -\log(1-u)$.

(a) transformation.

step 1: $u = e^{-x}$

step 2: $\frac{du}{dx} = -e^{-x}$

step 3: $f_X(x) = \frac{1}{1-0} e^{-x} = e^{-x} \sim \text{Exp}(1)$

(b) transformation

step 1: $u = 1 - e^{-y}$

step 2: $\frac{du}{dy} = e^{-y}$

step 3: $f_Y(y) = \frac{1}{1-0} \cdot e^{-y} = e^{-y} \sim \text{Exp}(1)$

#4 (a) Given $X \sim \text{Gamma}(\alpha_1, \beta)$, $Y \sim \text{Gamma}(\alpha_2, \beta)$
 $U = X + Y$

The MGF of U : $M_U(t) = M_X(t) \cdot M_Y(t) = \left(\frac{1}{1-\beta t}\right)^{\alpha_1} \cdot \left(\frac{1}{1-\beta t}\right)^{\alpha_2} = \left(\frac{1}{1-\beta t}\right)^{\alpha_1 + \alpha_2}$

This is the MGF for $\text{Gamma}(\alpha_1 + \alpha_2, \beta)$

Therefore: $U \sim \text{Gamma}(\alpha_1 + \alpha_2, \beta)$

(b) Find the pdf of $\frac{X+Y}{\beta}$

Let the R.V. $T = \frac{X+Y}{\beta} = \frac{U}{\beta}$ (U is a linear monotonic transformation of U , hence transformation is reasonable)

step 1: $u = \beta T$

step 2: $\frac{du}{dT} = \beta$

step 3: $f_U^T(u) \cdot \left| \frac{du}{dT} \right| = \frac{(\beta T)^{\alpha_1 + \alpha_2 - 1} e^{-\beta T / \beta}}{\Gamma(\alpha_1 + \alpha_2) \beta^{\alpha_1 + \alpha_2}} \cdot \beta = \frac{T^{\alpha_1 + \alpha_2 - 1} e^{-T}}{\Gamma(\alpha_1 + \alpha_2)} \sim \text{Gamma}(\alpha_1 + \alpha_2, 1)$

Therefore, $f_T(t) = \frac{t^{\alpha_1 + \alpha_2 - 1} e^{-t}}{\Gamma(\alpha_1 + \alpha_2)}$, which is $\text{Gamma}(\alpha_1 + \alpha_2, 1)$

(4)